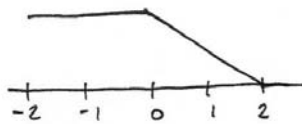


**4,4 Important Continuous Random Variables**

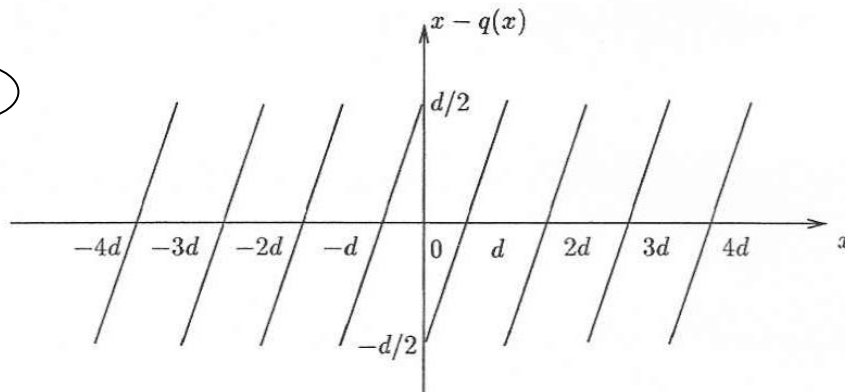
4.59  $P[|X| > x] = P[\{X > x\} \cup \{X < -x\}]$   
 $= P[X > x] + P[X < -x]$   
 $= 1 - F_X(x) + F_X(-x)$

$$F_X(x) = \begin{cases} 0 & x \leq -2 \\ \frac{x+2}{4} & -2 \leq x \leq 2 \\ 1 & x \geq 2 \end{cases}$$

$$P[|X| > x] = \begin{cases} 1 & x \leq 0 \\ 1 - \left(\frac{x+2}{4}\right) + \left(\frac{-x+2}{4}\right) = 1 - \frac{x}{2} & 0 \leq x \leq 2 \\ 0 & x \geq 2 \end{cases}$$



4.60



for  $-\frac{d}{2} < y < \frac{d}{2}$  the equation  $y = x - q(x)$  has 8 roots, thus from Eqn. 3.55:

$$f_Y(y) = \sum_{k=1}^8 \frac{f_X(x_k)}{\left. \frac{dy}{dx} \right|_{x=x_k}}$$

Since  $x - q(x)$  consists of piecewise linear unit-slope segments, we have that  $\left. \frac{dy}{dx} \right|_{x=x_k} = 1$  all  $x_k$ .

Thus

$$f_Y(y) = \sum_{k=1}^8 f_X(x_k) = \sum_{k=1}^8 \frac{1}{8d} = \frac{1}{d}$$

for  $-\frac{d}{2} < y < \frac{d}{2}$  ✓

5.29

a) For  $0 \leq y_0 \leq x_0$  we integrate along the strip indicated below.

b) The marginal cdf's are obtained by taking the appropriate limits of the joint cdf:

$$F_X(x_0) = \lim_{y_0 \rightarrow \infty} F_{XY}(x_0, y_0) = F_{XY}(x_0, x_0) = 1 - 2e^{-x_0} + e^{-2x_0}$$

$$F_Y(y_0) = \lim_{x_0 \rightarrow \infty} F_{XY}(x_0, y_0) = 1 - e^{-2y_0}$$

5.30

$$f_X(x) = \int_0^{\infty} x e^{-x} e^{-xy} dy = x e^{-x} \left( \frac{-1}{x} e^{-xy} \right)_0^{\infty} = e^{-x}$$

$$f_Y(y) = \int_0^{\infty} x e^{-x(1+y)} dx = \frac{e^{-x(1+y)}((1+y)x - 1)}{(1+y)^2} \Big|_0^{\infty}$$

$$= \frac{1}{(1+y)^2}$$

4.105

$$\begin{aligned}
 E[X] &= \left. \frac{1}{j} \frac{d}{d\omega} e^{jm\omega - \sigma^2\omega^2/2} \right|_{\omega=0} \\
 &= \left. \frac{1}{j} (jm - \sigma^2\omega) e^{jm\omega - \sigma^2\omega^2/2} \right|_{\omega=0} \\
 &= m \\
 E[X^2] &= \left. \frac{1}{j^2} \frac{d^2}{d\omega^2} e^{jm\omega - \sigma^2\omega^2/2} \right|_{\omega=0} \\
 &= \left. \frac{1}{j^2} \left[ -\sigma^2 e^{jm\omega - \sigma^2\omega^2/2} + (jm - \sigma^2\omega)^2 e^{jm\omega - \sigma^2\omega^2/2} \right] \right|_{\omega=0} \\
 &= \frac{1}{j^2} [-\sigma^2 + j^2 m^2] = \sigma^2 + m^2 \\
 \text{VAR}[X] &= E[X^2] - \mathcal{E}[X]^2 = \sigma^2
 \end{aligned}$$

4.106

$$\begin{aligned}
 \Phi_Y(\omega) &= E[e^{j\omega Y}] = E[e^{j\omega(aX+b)}] \\
 &= E[e^{j\omega aX}] e^{j\omega b} \\
 &= e^{j\omega b} \Phi_X(a\omega) \\
 &= e^{j\omega b} e^{j\omega m - \frac{\sigma^2\omega^2}{2}} \Big|_{\omega=a\omega} \\
 &= e^{j\omega b} e^{j\omega am - \frac{a^2\sigma^2\omega^2}{2}}
 \end{aligned}$$

characteristic fun for Gaussian RV  
 with mean  $a\omega + b$   
 and variance  $a^2\sigma^2$

### 5.9 Pairs of Jointly Gaussian Random Variables

5.110

$$f_{XY}(x,y) = \frac{e^{-(2x^2 + y^2/2)}}{2\pi c}$$

The sol'n involves matching the coefficients of the polynomial in the exponent of the Gaussian pdf.

$$\text{coeff. of } x^2 \Rightarrow \frac{1}{2(1-\rho^2)\sigma_1^2} = 2.$$

$$\text{coeff. of } y^2 \Rightarrow \frac{1}{2(1-\rho^2)\sigma_2^2} = \frac{1}{2}$$

$$\text{coeff. of } xy \Rightarrow \frac{-2\rho}{2(1-\rho^2)\sigma_1\sigma_2} = 0 \quad \therefore \rho = 0$$

$$\sigma_1^2 = \frac{1}{2(1-0^2) \cdot 2} = \frac{1}{4}$$

$$\sigma_2^2 = \frac{1}{2(1-0^2) \cdot \frac{1}{2}} = 1 //$$

$$\therefore \text{COV}(X,Y) = \rho \sigma_1 \sigma_2 = 0$$

$$\text{VAR}[X] = \frac{1}{4}$$

$$\text{VAR}[Y] = 1$$

**5.6 Joint Moments and Expected Value of a Function of Two Random Variables**

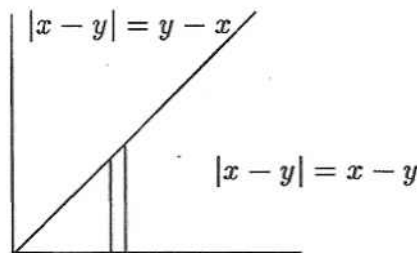
5.56

4.59 a)  $\mathcal{E}[(X + Y)^2] = \mathcal{E}[X^2 + 2XY + Y^2] = \mathcal{E}[X^2] + 2\mathcal{E}[XY] + \mathcal{E}[Y^2]$

b) 
$$\begin{aligned} \text{VAR}[X + Y] &= \mathcal{E}[(X + Y)^2] - \mathcal{E}[X + Y]^2 \\ &= \mathcal{E}[X^2] + 2\mathcal{E}[XY] + \mathcal{E}[Y^2] - \mathcal{E}[X]^2 \\ &\quad - 2\mathcal{E}[X]\mathcal{E}[Y] - \mathcal{E}[Y]^2 \\ &= \text{VAR}[X] + \text{VAR}[Y] + 2[\mathcal{E}[XY] - \mathcal{E}[X]\mathcal{E}[Y]] \end{aligned}$$

c)  $\text{VAR}[X + Y] = \text{VAR}[X] + \text{VAR}[Y]$  if  $\mathcal{E}[XY] = \mathcal{E}[X]\mathcal{E}[Y]$  that is, if  $X$  and  $Y$  are uncorrelated.

5.57



$$\begin{aligned} \mathcal{E}[|X - Y|] &= \int_0^\infty \int_0^\infty 2|x - y|e^{-(x+y)} dx dy \\ &= 2 \int_0^\infty \int_0^x (x - y)e^{-x}e^{-y} dy dx \\ &= 2 \int_0^\infty e^{-x} [x(1 - e^{-x}) - \underbrace{\int_0^x ye^{-y} dy}_{1 - (1+x)e^{-x}}] dx \\ &= 2 \int_0^\infty (xe^{-x} + e^{-2x} - e^{-x}) dx \\ &= 2 \left[ 1 + \frac{1}{2} - 1 \right] = 1 + \frac{4}{3} - 2 = \frac{1}{3} \end{aligned}$$

5.58

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$$\begin{aligned} \mathcal{E}[X^2 e^Y] &= \mathcal{E}[X^2] \mathcal{E}[e^Y] = 1 \cdot \mathcal{E}[e^Y] = \frac{1}{3} (e^3 - 1) \\ \mathcal{E}[X^2 Y] &= \mathcal{E}[X^2] \mathcal{E}[Y] = 1(1) = 1 \end{aligned}$$


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$$\frac{1}{3} \int_0^3 e^y dy = \frac{1}{3} e^y \Big|_0^3 = \frac{1}{3} (e^3 - 1)$$


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